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**Research and Development at U.S. Research Universities: An
Analysis of Scope Economies**

By

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Abstract:

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Keywords: R&D, university, patent, scope, complementarity, scale.

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Research and Development at U.S. Research Universities: An Analysis of Scope Economies

This work investigates the presence and sources of economies of scope in R&D at U.S. research universities. The analysis evaluates the tradeoffs and synergies arising between traditional university research outputs (articles and doctorates) and academic patents. We propose a new measure of economies of scope based on a primal representation of the underlying technology. We derive a decomposition of economies of scope which identifies its sources (e.g., complementarity effects and scale effects). Non-parametric estimates of scope economies using R&D input and output data from 92 research universities show significant economies of scope between articles and patents, but modest complementarities. (JEL O3, O31, O33, C6, L31)

Research and development (R&D) are fundamental to technological progress and economic growth. Because universities are dedicated to the production and dissemination of new knowledge and new technologies, university spillovers and their effects on economic growth have been the subject of much interest (e.g., Adam Jaffe, 1989; Rebecca Henderson et al., 1998; Bronwyn H. Hall et al., 2003b; Lee Branstetter, 2003; Suzanne Scotchmer, 2004). In the early 1980s, changes in federal policies, starting with the Bayh-Dole Act, made it easier for U.S. universities to retain the property rights to inventions obtained from federally funded research. This broad institutional change in intellectual property rights, combined with recent tightening in state and federal budgets, have helped to increase university efforts to secure both research sponsorship and intellectual property right royalties from the private sector.

Over the last ten years, academic patenting in the U.S. has increased sharply (Jeremy D. Foltz et al., 2005). University tech transfer offices, many recently established, intensified their

efforts to secure property rights to new knowledge and to transfer their research findings to the private sector through licensing arrangements, start-ups, and other remunerative arrangements. These efforts have raised a wide range of questions about the changing role of public and private research universities in the economy, society, and the pursuit of knowledge (e.g., Pierre Azoulay et al., 2004; Lee Branstetter, 2003; Bronwyn H. Hall et al., 2003b; Rebecca Henderson et al., 1998; Richard A. Jensen and Marie C. Thursby, 2001; Bhaven Sampat et al., 2003). One key issue is the existence of possible synergies between patenting and more traditional university outputs; i.e. the existence and nature of economies of scope within research universities.

Following the pioneering work by William J. Baumol et al. (1982), economies of scope measure the benefit for a firm to produce multiple outputs. Measuring such benefits for universities has provided useful insights into their organizational structure (e.g., Elchanan Cohn, et al., 1989; Hans De Groot et al., 1991; G. Thomas Sav, 2004). But, does the diversification of research universities into patenting activities generate significant synergies? In principle, university patenting and private-public partnering activities can help research universities become more effective in stimulating innovations (e.g., Hall et al., 2003b). However, at this point, the nature and magnitude of these benefits remain unclear. How large are these benefits? And how are they distributed among universities of different sizes or different types (e.g., private versus public universities)?

This paper investigates the presence and sources of economies of scope in R&D production at U.S. research universities. The analysis addresses the following issue in the literature on academic patenting: whether synergies arise between traditional university research outputs (articles and doctorates) and the more recent and burgeoning output of academic patents. Framing the empirical analysis requires a theoretical exposition of the concept of economies of

scope that deepens our understanding of this phenomenon in ways that are relevant not only to R&D processes but also to the many other economic contexts where scope economies may arise. Overall, our paper makes three methodological contributions.

First, the conventional approach to measuring economies of scope typically involves analyzing complete specialization among outputs (see Baumol et al., 1982). This is relevant in the evaluation of mergers and acquisitions when firms are deciding on whether to produce jointly distinctive outputs or to spin off separate operations. However, universities are rarely completely specialized. On that basis, we develop an analysis that allows for partial specialization between production processes. Allowing for partial specialization permits a search for economies of scope across a more nuanced range of possible outcomes than is typically depicted in previous economies of scope studies.

The second methodological contribution of this paper is to develop and apply a primal approach to economies of scope. The approach relies on David G. Luenberger's (1995) shortage function as a representation of the underlying technology. This primal approach is especially useful in measuring scope economies in contexts where input cost data are difficult to obtain or fraught with measurement problems, such as in our empirical study with respect to certain input prices that shape university research and teaching outputs.

The third and perhaps most far-reaching contribution of this paper is its decomposition of economies of scope into three measures: complementarities between outputs, economies of scale in multiple outputs (along different degrees of specialization), and a convexity component. This decomposition provides a clear picture of the basis for scope economy outcomes in production of multiple outputs, because it permits identification of whether the scale of operation and/or the complementarity of outputs are driving scope benefits. This decomposition extends the

contributions of Paul Milgrom and John Roberts (1990) by making identification of complementarity and other sources of scope economies both more tractable and intuitive. Our empirical analysis of U.S. research universities illustrates how both scale and complementarity can drive scope outcomes for economic entities as their size and type change.

Finally, our scope analysis is applied to U.S. research universities. This provides evidence on the synergies that exist between patenting and more traditional academic research outputs. R&D input and output data from 92 top-tier U.S. research universities are used to estimate economies of scope for public and private universities of different sizes. The estimates are obtained using a non-parametric representation of the underlying technology, and control for quality of article and patent outputs. The empirical results show that significant variations exist in the magnitude and sources of economies of scope across U.S. universities. Indeed, while scope economies are evident in most of the sample, evidence for strong complementarity among these research activities is more limited.

The organization of the paper is as follows. Section 2 provides the basic model and a characterization of firms going from integrated, to mildly specialized, to fully specialized output choices. Section 3 proposes a new primal measure of scope economies based on Luenberger's (1995) shortage function. Section 4 presents a decomposition of scope economies which identifies the potential sources of scope benefits among outputs. Section 5 discusses the dataset on university research outputs and inputs. Using a non-parametric approach, section 6 applies our proposed methodology to estimate scope economies and their sources across a spectrum of specialization for U.S. research universities. Section 7 reports the empirical results, and section 8 concludes.

I. The Model

Consider a firm facing a production process producing m outputs using n inputs, where $\mathbf{y} = (y_1, \dots, y_m) \in \mathbf{R}_+^m$ is the vector of outputs, and $\mathbf{x} \in \mathbf{R}_+^n$ is the vector of inputs. Using the netput notation (where inputs are negative and outputs are positive), the netputs are $\mathbf{z} \equiv (-\mathbf{x}, \mathbf{y})$. The technology is represented by the production possibility set $F \subset \mathbf{R}_-^n \times \mathbf{R}_+^m$, where $\mathbf{z} \equiv (-\mathbf{x}, \mathbf{y}) \in F$ means that outputs \mathbf{y} can be produced from inputs \mathbf{x} . Throughout the paper, we assume that the set F is closed and with a non-empty interior. We want to investigate under what conditions the multiproduct firm would gain (or lose) from reorganizing its production activities in a more specialized way. The reorganization involves breaking up the firm into K specialized firms, $2 \leq K \leq m$. Given the output index $I = \{1, \dots, m\}$, consider its partition $I = \{I_{A1}, I_{A2}, \dots, I_{AK}, I_B\}$, where $I_A = \{I_{A1}, I_{A2}, \dots, I_{AK}\}$, I_{Ak} being the set of outputs that the k -th firm is specializing in, $k = 1, \dots, K$, while I_B being the set of outputs that no particular firm specializes in. Let $\mathbf{y}^k = (y_1^k, \dots, y_m^k)$ denote the outputs produced by the k -th specialized firm, $k = 1, \dots, K$.

Our analysis of the economics of specialization has two objectives in mind. First, we explore what happens under alternative specialization schemes holding total output constant. This requires selecting the outputs of specialized firms such that $\sum_{k=1}^K \mathbf{y}^k = \mathbf{y}$, where the K specialized firms produce the same aggregate output vector \mathbf{y} as the original firm. Second, we want to allow various degrees of specialization, going from “mild specialization” to “complete specialization”. In this context, given $\mathbf{y} = (y_1, \dots, y_m)$, consider the following patterns of specialization for the k -th firm

$$\begin{aligned}
 (1a) \quad & y_i^k = \beta y_i, \text{ if } i \in I_{Ak}, \\
 (1b) \quad & = y_i (1-\beta)/(K-1), \text{ if } i \in I_{Ak}, k' \neq k, \\
 (1c) \quad & = y_i/K, \text{ if } i \in I_B,
 \end{aligned}$$

for some β , $1/K < \beta \leq 1$, $k = 1, \dots, K$. This represents a reorganization of the original firm into K firms toward greater specialization, where the k -th firm becomes more specialized in the production of outputs in the sets I_{Ak} , $k = 1, \dots, K$.

Note that the specification (1a)-(1c) always satisfies $\sum_{k=1}^K y_i^k = y_i$, $i = 1, \dots, m$. This guarantees that the same aggregate outputs are being produced before and after the firm reorganization. The parameter β in (1a) represents the proportion of the original outputs $\{y_i: i \in I_{Ak}\}$ produced by the k -th firm. And from (1b), $(1-\beta)$ represents the proportion of the original outputs $\{y_i: i \in I_{Ak'}, k' \neq k\}$ produced by the k -th firm. When $\beta = 1$, this means that the k -th firm produces the same quantities $\{y_i: i \in I_{Ak}\}$ as the original firm and that such outputs are produced only by the k -th firm. In this case, the k -th firm is completely specialized in the production of the outputs in the set I_{Ak} (and it produces none of the other outputs in the sets I_A). Alternatively, when $1/K < \beta < 1$, we allow for partial specialization. For example, if $K = 2$ and $\beta = 0.9$, then the first firm (corresponding to $k = 1$) produces 90% of the quantities $\{y_i: i \in I_{A1}\}$ produced by the original firm, while the second firm (corresponding to $k = 2$) produces the remaining 10%. And the second firm ($k = 2$) produces 90% of the quantities $\{y_i: i \in I_{A2}\}$ produced by the original firm, while the first firm ($k = 1$) produces the remaining 10%. Finally, note that (1c) allocates the outputs in the set I_B equally among the K specialized firms. This simply reflects that the outputs in I_B are not involved in the patterns of specialization as the firm reorganizes.

Equations (1a)-(1c) include as a special case the situation where $\beta = 1$ and $I_B = \emptyset$. This is the case of complete specialization (e.g., as investigated by Baumol et al., 1982 based on a cost function). As such, our approach extends previous analyses in two directions. First it allows for specialization in a subset of outputs (when $I_B \neq \emptyset$). This can become relevant in the economics of specialization when $2 \leq K < m$, i.e., when the number of specialized firms is less than the

number of outputs. Second, as noted above, it allows for partial specialization in the outputs of the set I_A (with $1/K < \beta < 1$). This is relevant when the K firms want to explore the economics of becoming more specialized (thus deemphasizing the production of some of their outputs) but without a complete shutdown of some of their production lines.

II. Economies of Scope

To investigate the economics of specialization, we need to rely on measures that can be meaningfully added across firms. This is the case of the cost function which has provided the standard basis for measuring economies of scope. In this context, Baumol et al. (1982) have defined economies of scope (diseconomies of scope) as situations where it is less costly (more costly) to produce the aggregate outputs y from an integrated firm as compared to specialized firms. This has stimulated empirical analyses of the benefit (or cost) of producing from an integrated multi-output firm. However, the cost function requires that all inputs be market goods with observable prices. There are situations where some inputs have prices that are not observable or that do not reflect their marginal contribution to the production process. An example in higher education includes Ph.D. students: their cost to a university can differ significantly from their marginal contribution to university research productivity. Under such scenarios, the use of the cost function becomes problematic. But under the convexity assumption, it is well known that the cost function is dual to the underlying technology. This means that there are alternative ways of measuring economies of scope directly from the production technology. Like the cost function, this requires using a measurement of the production technology that can be meaningfully added across firms. A measurement that satisfies this property is Luenberger's

shortage function, which we use below in our analysis of the scope economies associated with integrated production.

Following Luenberger (1995), letting $\mathbf{g} \in \mathbf{R}_+^{n+m} - \{0\}$ be some reference netput bundle, define the shortage function:

$$(2) \quad \sigma(\mathbf{z}, \mathbf{g}) = \min_{\gamma} \{ \gamma : (\mathbf{z} - \gamma \mathbf{g}) \in F \} \text{ if } (\mathbf{z} - \gamma \mathbf{g}) \in F \text{ for some scalar } \gamma, \\ = +\infty \text{ otherwise.}$$

The shortage function $\sigma(\mathbf{z}, \mathbf{g})$ in (2) measures how far the point \mathbf{z} is from the frontier of technology, expressed in units of the reference bundle \mathbf{g} . To illustrate, consider the case where $\mathbf{g} = (0, \dots, 0, 1)$. Then, the shortage function is $\sigma(\mathbf{z}, \mathbf{g}) = \min_{\gamma} \{ \gamma : (z_1, \dots, z_{m-1}, z_m - \gamma) \in F \} = z_m - f(z_1, \dots, z_{m-1})$, where $f(z_1, \dots, z_{m-1})$ is a (multi-output) production frontier, and feasibility implies that $z_m \leq f(z_1, \dots, z_{m-1})$. Under differentiability, this implies that $\partial \sigma / \partial z_i = -\partial f / \partial z_i$, i.e. that the marginal shortage $\partial \sigma / \partial z_i$ is the negative of the marginal product $\partial f / \partial z_i$ with respect to the i -th netput, $i = 1, \dots, m-1$. Note that, given a reference bundle \mathbf{g} , the shortage function can be meaningfully added across firms. As such, the shortage function provides a convenient basis for analyzing scope issues and the benefit/cost of specialization.¹

Starting from a firm using netputs $\mathbf{z} \equiv (-\mathbf{x}, \mathbf{y})$, we analyze whether there are any benefits from reorganizing its production activities according to equation (1), where $\mathbf{y}^k \in \mathbf{R}_+^m$ is produced by the k -th specialized firm, $k = 1, \dots, K$, with $\mathbf{y} = \sum_{k=1}^K \mathbf{y}^k$. If the k -th firm uses inputs \mathbf{x}^k , the shortage function associated with $(-\mathbf{x}^k, \mathbf{y}^k)$ is $\sigma(-\mathbf{x}^k, \mathbf{y}^k, \mathbf{g})$. In a way similar to (1c), consider the case where inputs \mathbf{x} are equally divided between the K firms, with $\mathbf{x}^k = \mathbf{x}/K$, $k = 1, \dots, K$.

Definition 1: Given equations (1), economies of scope (diseconomies of scope) with respect to the partition $I = \{I_{A1}, \dots, I_{AK}, I_B\}$ in the production of outputs \mathbf{y} are said to exist if

$$(3) \quad S(\beta, I_{A1}, \dots, I_{AK}, I_B, \mathbf{z}, \mathbf{g}) \equiv \sum_{k=1}^K \sigma(-\mathbf{x}/K, \mathbf{y}^k, \mathbf{g}) - \sigma(\mathbf{z}, \mathbf{g}) > (<) 0,$$

Note that $\sum_{k=1}^K \sigma(-\mathbf{x}/K, \mathbf{y}^k, \mathbf{g})$ can be interpreted as the smallest distance to the technology frontier (as measured by the number of units of the reference bundle \mathbf{g}) when the aggregate netputs $\mathbf{z} = (-\mathbf{x}, \mathbf{y})$ are produced by K specialized firms: $(-\mathbf{x}/K, \mathbf{y}^k)$, $k = 1, \dots, K$. Thus, equation (3) compares the distance to the technology frontier producing \mathbf{y} from an integrated firm versus specialized firms.

To help interpret (3), consider the case where netputs are market goods with prices $\mathbf{p} \in \mathbf{R}_{++}^{n+m}$. Then, starting from the aggregate netput \mathbf{z} and under technical efficiency, $\pi_a = \mathbf{p} \cdot [\mathbf{z} - \sigma(-\mathbf{x}, \mathbf{y}, \mathbf{g}) \mathbf{g}]$ is the profit for the integrated firm, while $\pi_s = \mathbf{p} \cdot [\sum_{k=1}^K \mathbf{z}^k - \sigma(-\mathbf{x}/K, \mathbf{y}^k, \mathbf{g}) \mathbf{g}]$ is the aggregate profit for the K specialized firms, where $\mathbf{z}^k = (-\mathbf{x}/K, \mathbf{y}^k)$, and \mathbf{y}^k satisfies (1), $k = 1, \dots, K$. It follows that the difference in profit is

$$\pi_a - \pi_s = [\sum_{k=1}^K \sigma(-\mathbf{x}/K, \mathbf{y}^k, \mathbf{g}) - \sigma(-\mathbf{x}, \mathbf{y}, \mathbf{g})] \mathbf{p} \cdot \mathbf{g},$$

where $(\pi_a - \pi_s)$ measures the benefit of integrated production in a multiproduct firm. When positive, this difference reflects positive synergy among outputs. Given $\mathbf{p} \cdot \mathbf{g} > 0$, this makes it clear that $S(\beta, I_{A1}, \dots, I_{AK}, I_B, \mathbf{z}, \mathbf{g}) > 0$ in (3) corresponds to economies of scope, identifying the presence of synergies or positive externalities in the production process among the outputs in I_{Ak} , $k = 1, \dots, K$. Alternatively, diseconomies of scope exist (with $S(\beta, I_{A1}, \dots, I_{AK}, I_B, \mathbf{z}, \mathbf{g}) < 0$) if producing netputs \mathbf{z} from an integrated firm (as opposed to K specialized firms) reduces the benefits. This identifies the presence of negative externalities in the production process among the outputs in I_{Ak} , $k = 1, \dots, K$.

How does S in (3) compare with the traditional cost-based measure of scope proposed by Baumol et al. (1982)? They define economies of scope when $S' \equiv \sum_{k=1}^K C(\mathbf{r}, \mathbf{y}^k) - C(\mathbf{r}, \mathbf{y}) > 0$, where \mathbf{r} is the input price vector and $C(\mathbf{r}, \mathbf{y}) = \min_{\mathbf{x}} \{\mathbf{r} \cdot \mathbf{x} : (-\mathbf{x}, \mathbf{y}) \in F\}$ is the cost function. Consider the case where $\mathbf{g} = (\mathbf{g}_x, 0)$, and \mathbf{x} is the cost-minimizing input bundle under outputs \mathbf{y} : $\mathbf{x} \in \operatorname{argmin}_{\mathbf{x}} \{\mathbf{r} \cdot \mathbf{x}' : (-\mathbf{x}', \mathbf{y}) \in F\}$. Note that cost minimization implies that $C(\mathbf{r}, \mathbf{y}^k) \leq \mathbf{r} \cdot [\mathbf{x}/K - \sigma(-\mathbf{x}/K, \mathbf{y}^k, \mathbf{g}_x) \mathbf{g}_x]$. For given input prices \mathbf{r} , it follows that $S' = \sum_{k=1}^K C(\mathbf{r}, \mathbf{y}^k) - \mathbf{r} \cdot \mathbf{x} \leq \sum_{k=1}^K \mathbf{r} \cdot [\mathbf{x}/K - \sigma(-\mathbf{x}/K, \mathbf{y}^k, \mathbf{g}_x) \mathbf{g}_x] = (\mathbf{r} \cdot \mathbf{g}_x) S$. When input prices are normalized such that $\mathbf{r} \cdot \mathbf{g}_x = 1$, this implies that $S' \leq S$: the Baumol scope measure S' is a lower bound on S in (3). This reflects possible allocative inefficiencies when \mathbf{x}/K does not minimize the cost of producing \mathbf{y}^k , $k = 1, \dots, K$. Alternatively, if $\mathbf{r} \cdot \mathbf{g}_x = 1$ and \mathbf{x}/K did minimize the cost of producing each \mathbf{y}^k , then $S' = S$ and the two scope measures become identical. Of course, this is conditional on input prices \mathbf{r} . In situations where input prices are difficult to assess, then only the primal measure S in (3) remains empirically tractable.

III. A Decomposition of Economies of Scope

For simplicity, we focus our attention on the case of splitting the original firm (which produces the output vector \mathbf{y}) into two firms ($K = 2$).² Then, with the partition $I = \{I_{A1}, I_{A2}, I_B\}$, the first firm ($k = 1$) specializes in the outputs in I_{A1} , the second firm ($k = 2$) specializes in the outputs in I_{A2} , and $\mathbf{y} = (\mathbf{y}_{A1}, \mathbf{y}_{A2}, \mathbf{y}_B)$, where $\mathbf{y}_{A1} = \{y_i : i \in I_{A1}\}$, $\mathbf{y}_{A2} = \{y_i : i \in I_{A2}\}$, $\mathbf{y}_A = (\mathbf{y}_{A1}, \mathbf{y}_{A2})$, and $\mathbf{y}_B = \{y_i : i \in I_B\}$ are the remaining outputs. From equations (1), it follows that $\mathbf{y}^1 = (\beta \mathbf{y}_{A1}, (1-\beta) \mathbf{y}_{A2}, \frac{1}{2} \mathbf{y}_B)$, and $\mathbf{y}^2 = ((1-\beta) \mathbf{y}_{A1}, \beta \mathbf{y}_{A2}, \frac{1}{2} \mathbf{y}_B)$.

A useful decomposition of S in (3) is presented next. See the proof in Appendix A.

Proposition 1: Assume that the shortage function $\sigma(\mathbf{z}, \mathbf{g})$ is continuous in \mathbf{z} and differentiable almost everywhere in $\mathbf{y} \in \mathbf{R}_+^m$. Under equations (1) with $K = 2$, there are economies of scope in the production of outputs $\mathbf{y} = (\mathbf{y}_{A1}, \mathbf{y}_{A2}, \mathbf{y}_B) \in \mathbf{R}_{++}^m$ if and only if

$$(4) \quad \begin{aligned} S(\beta, I_{A1}, I_{A2}, I_B, \mathbf{z}, \mathbf{g}) &\equiv S_C(\beta, I_{A1}, I_{A2}, I_B, \mathbf{z}, \mathbf{g}) + S_R(\beta, I_{A1}, I_{A2}, I_B, \mathbf{z}, \mathbf{g}) \\ &+ S_I(\beta, I_{A1}, I_{A2}, I_B, \mathbf{z}, \mathbf{g}) > 0, \end{aligned}$$

where

$$(5a) \quad \begin{aligned} S_C(\beta, I_{A1}, I_{A2}, I_B, \mathbf{z}, \mathbf{g}) &\equiv - \int_{(1-\beta)y_{A2}}^{\beta y_{A2}} [\partial \sigma / \partial \gamma (-\frac{1}{2} \mathbf{x}, \beta \mathbf{y}_{A1}, \gamma, \frac{1}{2} \mathbf{y}_B, \mathbf{g}) \\ &- \partial \sigma / \partial \gamma (-\frac{1}{2} \mathbf{x}, (1-\beta) \mathbf{y}_{A1}, \gamma, \frac{1}{2} \mathbf{y}_B, \mathbf{g})] d\gamma, \end{aligned}$$

$$(5b) \quad S_R(\beta, I_{A1}, I_{A2}, I_B, \mathbf{z}, \mathbf{g}) \equiv 2 \sigma(\frac{1}{2} \mathbf{z}, \mathbf{g}) - \sigma(\mathbf{z}, \mathbf{g}),$$

$$(5c) \quad \begin{aligned} S_I(\beta, I_{A1}, I_{A2}, I_B, \mathbf{z}, \mathbf{g}) &\equiv \sigma(-\frac{1}{2} \mathbf{x}, \beta \mathbf{y}_A, \frac{1}{2} \mathbf{y}_B, \mathbf{g}) + \sigma(-\frac{1}{2} \mathbf{x}, (1-\beta) \mathbf{y}_A, \frac{1}{2} \mathbf{y}_B, \mathbf{g}) \\ &- 2 \sigma(\frac{1}{2} \mathbf{z}, \mathbf{g}). \end{aligned}$$

Proposition 1 gives a necessary and sufficient condition for economies of scope in the production of outputs \mathbf{y} . Equation (4) decomposes the scope measure $S(\beta, I_{A1}, I_{A2}, I_B, \mathbf{z}, \mathbf{g})$ in (3) into three additive terms: $S_C(\beta, I_{A1}, I_{A2}, I_B, \mathbf{z}, \mathbf{g})$ given in (5a), $S_R(\beta, I_{A1}, I_{A2}, I_B, \mathbf{z}, \mathbf{g})$ given in (5b), and $S_I(\beta, I_{A1}, I_{A2}, I_B, \mathbf{z}, \mathbf{g})$ given in (5c).

The term S_C in (5a) depends on how \mathbf{y}_{A1} affects the marginal shortage of \mathbf{y}_{A2} . As illustrated in section 3, marginal shortage can be interpreted as the negative of the marginal product. With this interpretation in mind, given $\beta \in (0.5, 1]$, we define complementarities between \mathbf{y}_{A1} and \mathbf{y}_{A2} at point \mathbf{y} as any situation where the shortage function $\sigma(\mathbf{z}, \mathbf{g})$ satisfies $[\partial \sigma / \partial \mathbf{y}_{A2}(-\frac{1}{2} \mathbf{x}, \beta \mathbf{y}_{A1}, \gamma \mathbf{y}_{A2}, \frac{1}{2} \mathbf{y}_B, \mathbf{g}) - \partial \sigma / \partial \mathbf{y}_{A2}(-\frac{1}{2} \mathbf{x}, (1-\beta) \mathbf{y}_{A1}, \gamma \mathbf{y}_{A2}, \frac{1}{2} \mathbf{y}_B, \mathbf{g})] \leq 0$ for all $\gamma \in [0, 1]$, with the inequality being strict over a set of nonzero measure. Then, it is clear from (5a) that

$S_C > 0$ if the shortage function exhibits complementarities between \mathbf{y}_{A1} and \mathbf{y}_{A2} . Thus, the term S_C can be interpreted as reflecting the role of complementarities between \mathbf{y}_{A1} and \mathbf{y}_{A2} in economies of scope.

Note that when, the shortage function $\sigma(\mathbf{z}, \mathbf{g})$ is twice differentiable in $\mathbf{y} \in \mathbf{R}_+^m$, then S_C in (5a) can be alternatively written as

$$(5a') \quad S_C \equiv - \int_{(1-\beta)\mathbf{y}_{A2}}^{\beta\mathbf{y}_{A2}} \int_{(1-\beta)\mathbf{y}_{A1}}^{\beta\mathbf{y}_{A1}} \partial^2 \sigma / \partial \gamma_1 \partial \gamma_2 (-\frac{1}{2} \mathbf{x}, \gamma_1, \gamma_2, \frac{1}{2} \mathbf{y}_B, \mathbf{g}) \, d\gamma_1 \, d\gamma_2.$$

When $\beta \in (0.5, 1]$, equation (5a') makes it clear that the sign of S_C depends on the sign of $\partial^2 \sigma / \partial \mathbf{y}_{A1} \partial \mathbf{y}_{A2}$. This shows that, under twice differentiability, complementarities can be defined as any situation where $\partial^2 \sigma / \partial \mathbf{y}_{A1} \partial \mathbf{y}_{A2} (-\frac{1}{2} \mathbf{x}, \gamma_1 \mathbf{y}_{A1}, \gamma_2 \mathbf{y}_{A2}, \frac{1}{2} \mathbf{y}_B, \mathbf{g}) \leq 0$ for all $\gamma_i \in [0, 1]$, $i = 1, 2$, with the inequality being strict over a set of nonzero measure. Recall that the term $\partial \sigma / \partial y_i$ can be interpreted as the negative of the marginal product with respect to y_i . Thus, when $\partial^2 \sigma / \partial \mathbf{y}_{A1} \partial \mathbf{y}_{A2} < 0$, complementarities mean that \mathbf{y}_{A1} has positive effects on the marginal product of \mathbf{y}_{A2} , implying positive synergies between \mathbf{y}_{A1} and \mathbf{y}_{A2} (see Baumol et al., 1982; Milgrom and Roberts, 1990).

To interpret the term S_R in (5b), using lemma 1 in Appendix A, note that $2 \sigma(\frac{1}{2} \mathbf{z}, \mathbf{g}) <, =,$ or $> \sigma(\mathbf{z}, \mathbf{g})$ under decreasing return to scale (DRTS), constant return to scale (CRTS), or increasing return to scale (IRTS), respectively. It follows that

$$(5b') \quad S_R(\beta, I_{A1}, I_{A2}, I_B, \mathbf{z}, \mathbf{g}) \begin{cases} < \\ = \\ > \end{cases} 0 \text{ under } \begin{cases} \text{DRTS} \\ \text{CRTS} \\ \text{IRTS} \end{cases}.$$

Equation (5b') implies that S_R vanishes under CRTS, but is positive (negative) under IRTS (DRTS). Thus, the term S_R can be interpreted as capturing scale effects generated as the output vector \mathbf{y} is produced by more specialized firms. Also, equation (5b') shows that $S_R \geq 0$ under non-decreasing returns to scale.

Finally, the term $S_V(\beta, I_{A1}, I_{A2}, I_B, \mathbf{z}, \mathbf{g})$ in (5c) reflects the effect of convexity. From lemma 2 in Appendix A, if the technology F is convex, the shortage function $\sigma(\mathbf{z}, \mathbf{g})$ is convex in \mathbf{z} and satisfies $\sigma(\theta \mathbf{z} + (1-\theta) \mathbf{z}', \mathbf{g}) \leq \theta \sigma(\mathbf{z}, \mathbf{g}) + (1-\theta) \sigma(\mathbf{z}', \mathbf{g})$ for any $\theta \in [0, 1]$ and any \mathbf{z} and \mathbf{z}' . Choosing $\theta = 1/2$, it follows that $S_V(\beta, I_{A1}, I_{A2}, I_B, \mathbf{z}, \mathbf{g}) \geq 0$ under a convex technology. In other words, a convex technology is sufficient to imply that $S_V \geq 0$. In addition, note that $S_V = 0$ when $\beta = 0.5$. Thus, under a convex technology, one can expect S_V to increase with the degree of specialization $\beta \in [0.5, 1]$.

The decomposition provided in Proposition 1 indicates that there can be multiple sources of economies of scope. Identifying the role played by each source appears useful as it can provide useful insights into the economics of specialization. This is illustrated next in an application to U.S. universities.

IV. Data

The dataset combines information on research inputs and outputs in the sciences and engineering for 92 US universities, including 61 public universities and 31 private universities for the period of 1995-1998. This dataset contains for all 92 universities the following data:

1) Total patent counts and patent citations from all science and engineering fields (U.S. Patent Office, 2004; and Hall et al., 2003a),

2) Article counts and citations from all science and engineering fields (ISI Web of Science, 2004),

3) Total number of doctorates and bachelor degrees granted in the sciences as well as the number of graduate students, faculty, and post-docs (National Science Foundation, 2004).

Further details on the sources of the data and key choices in the construction of the dataset can be found in Appendix B. One key aspect of the dataset warrants discussion here. The dataset focuses on scientific inputs and outputs, reflecting our interest in studying economies of scope between university research and university patents. Thus, our measures of scientific inputs and outputs are appropriate to investigate the possible tradeoff that exists between university research outputs and university patents (which are almost entirely produced by the sciences).

In order to proceed with the empirical analysis, we need a representation of the university production process. In the case of student training, we measure undergraduate bachelor's degrees in the sciences as university outputs. However, graduate students can be both inputs and outputs: they are outputs of the university educational function; but they are also inputs into the research process (especially through their theses and dissertations). To account for this dual function of graduate students, we assume that they are outputs, except in their final year when they are treated as inputs into the university research process. Since there is typically a one or two year delay between when research takes form as an article or patent and when a graduate student worked on it, we think that our assumption is a reasonable match with the output data we have. Thus, we measure continuing graduate students as outputs, and PhD's granted as inputs. In this context, universities are involved in the production of four outputs (journal articles, patents, trained undergraduate students, and trained graduate students) using three inputs (faculty, post-doctoral researchers, and PhD graduate students).

To account for quality differentials, quality adjustments are made on university output measures of patents and articles as well as input measure of faculty. Quality-adjusted output measures are obtained, where citations of articles and patents are used to control for quality of those two research outputs. Quality-adjusted input measure for faculty is obtained by multiplying

total faculty numbers and the university's average faculty salary (National Science Foundation, 2004). Science patent assignee and citation information were obtained from the NBER patent database (Hall et al., 2003a), while the Science Citation Index (ISI Web of Science, 2004) provided the science article and citation counts by year for each university. Patents are credited by application year rather than by grant date in order to measure them as close as possible to the date research efforts occurred. Quality adjustments were sought because in the case of research output, quality is likely to matter significantly to the implicit value of the research and also to the potential synergies between patents and articles. In the first case, highly cited articles and patents are likely to generate flows of additional research or licensing funds to the author or assignee, while in the latter research that gives rise, for example, to an article that is highly cited may also be more likely to generate a patent than would a larger number of un-cited articles. Empirically, studies of patent citations have shown that they provide a reasonable proxy for both the quality of a patent and knowledge spillovers from patents, because each time a new patent uses a piece of research from another patent it is obligated to cite the previous patent (Henderson et al., 1998). Article citations are also commonly used as measures of quality in studies of departmental or university quality (e.g., Amy Adams, 1998).

Using citations as a quality measure requires attending to the time dependency of the counts, namely the truncation problem associated with more recent articles or patents that may not have had time to generate many citations (Sampat et al., 2003). The quality adjustment measure used for each life science article/patent is the deviation from the average citation rate of an article/patent in the same broad class/category published in the same year. For example, a 1995 biochemistry article with 10 citations is compared to the average level of citations of all biochemistry articles produced in that year. For a given year, the average article within a

category has a citation rate of 1, with higher quality articles then having a measure greater than one and lower quality articles receiving a measure between zero and one. This relative citation approach minimizes a truncation bias that would be introduced by using an absolute citation count. Further details on the citation measure are provided in Appendix B.

Finally, we account for the fact that the production process for universities is dynamic: the process of scientific discovery is typically time-consuming. For example, lagged inputs can affect current outputs in the presence of production lags (e.g., it takes time for research to be published). And lagged outputs may affect current outputs in the presence of temporal synergies in production. This implies a need to incorporate dynamics in the representation of the underlying technology. This is done by specifying and estimating a multi-period production technology over a four-year period. Outputs for the current year are assumed to depend on inputs of the current year, but also on inputs and outputs from the three previous years. The effects of lagged quantities are captured by a weighted average of the corresponding quantities, with weight equal to 0.5 for lag one-year, 0.37 for lag two-years, and 0.13 for lag three-years. As a result, our dynamic production process is represented by eight outputs (four current outputs and four lagged outputs) and six inputs (three current inputs, and three lagged inputs). Our empirical investigation of economies of scope between university research outcomes (patents and articles) relies on data for 1995-1998 (the most recent years with complete data available).

V. Empirical Analysis

The shortage function described in equation (2) provides a generic representation of the frontier of technology. It can be estimated either using parametric methods (involving a parametric specification followed by an econometric estimation of the parameters) or non-

parametric methods. Below, we rely on a non-parametric approach for several reasons. First, it provides a flexible representation of the multi-output production frontier. Second, it does not require imposing a parametric structure on the problem. Third, when the number of netputs is large, it is not subject to collinearity problems. Finally, it does not require that each data point be on the frontier technology, which allows for possible technical inefficiencies (see Foltz et al., 2005).

Thus, we use input and output data to recover an estimate of the underlying multi-output production technology for universities. Again, this is done by representing the dynamic process of producing outputs (research articles, patents, undergraduate degrees granted, and graduate students (excluding final-year doctorates)) using a set of inputs (post-docs, doctorates in their final year of study, and faculty). And using the shortage function in (2), we have measurements of how far is each point from the production frontier.

Next, a nonparametric representation of economies of scope is investigated. To assess economies of scope between research articles and patents, we break up the original university into 2 specialized universities. We then identify the following partition I : I_{A1} = patents, I_{A2} = research articles, I_B = doctoral students in labs and bachelor degrees. This partition corresponds to scenarios of increasing specialization, where one university may specialize in patents while another in research publications. Note also that doctoral students in labs and bachelor degrees are included in the set of outputs that no particular firm specializes in. As shown in section 3, economies of scope (diseconomies of scope) with respect to the partition $I = \{I_{A1}, I_{A2}, I_B\}$ are defined as in equation (3).

Evaluating equation (3) requires the estimation of the shortage function under alternative scenarios. This requires first choosing a reference bundle g that will be the same under each

scenario. Our chosen reference bundle $\mathbf{g} = (g_1, \dots, g_{n+m})$ involves choosing $g_i = 1$ for current and lagged faculty input, $g_i = 0.00513$ for current and lagged post-docs, $g_i = 0.00335$ for current and lagged doctorates in their final year of study, and $g_i = 0$ otherwise. The numbers 0.00513 and 0.00335 are the ratios of post-docs per faculty, and of final-year doctorates per faculty, evaluated at sample means. The choice $g_i = 1$ for faculty means that our reference bundle can be interpreted as a typical input bundle associated with one faculty. Here faculty is measured in terms of (adjusted) faculty salary. With all shortage measurements being made in terms of units of this reference bundle \mathbf{g} , it follows that such measurements can be interpreted in terms of changes in faculty salaries, with proportional adjustments in post-docs and final-year doctorates. This provides a simple and logical measure of the distance from the frontier technology under alternative scenarios. A strength of this approach, as opposed to a dual counterpart (a cost function approach), is that the price information of some major inputs (e.g., post-docs, doctorates in their final year of study) are not required to assess economies of scope.

The non-parametric estimation of the technology and the associated shortage function is done as follows. Following Sidney Afriat (1972), Hal R. Varian (1984) and others, given a set of observations on T universities, $\mathbf{z}^t \equiv (-\mathbf{x}^t, \mathbf{y}^t)$, $t = 1, \dots, T$, a nonparametric representation of the technology under variable return to scale (VRTS) is $F^V = \{\mathbf{z}: \sum_{t=1}^T \lambda_t \mathbf{z}^t \geq \mathbf{z}, \sum_{t=1}^T \lambda_t = 1, \lambda_t \geq 0, t = 1, \dots, T\}$. In general, the set F^V is convex and satisfies $(-\mathbf{x}^t, \mathbf{y}^t) \in F^V$, $t = 1, \dots, T$. It does not require that all firms be technically efficient. Indeed, while technically efficient firms are necessarily located on the boundary of F^V , it allows for technically inefficient firms (located in the interior of F^V). Finally, it allows for increasing, constant, as well as decreasing return to scale. Then, given F^V , a nonparametric estimate of the shortage function under VRTS is

$$(6) \quad \sigma(-\mathbf{x}, \mathbf{y}, \mathbf{g}) = \min_{\gamma, \lambda} \{ \gamma: \sum_{t=1}^T \lambda_t \mathbf{z}^t \geq \mathbf{z} - \gamma \mathbf{g}, \sum_{t=1}^T \lambda_t = 1, \lambda_t \geq 0, t = 1, \dots, T \}.$$

This is standard linear programming problem.³ It can be solved for different values of $\mathbf{z} \equiv (-\mathbf{x}, \mathbf{y})$. For example, when evaluated at \mathbf{y}^k and $\mathbf{x}/2$ (as given in (1) where $\sum_{k=1}^2 \mathbf{y}^k = \mathbf{y}$), this yields $\sigma(\mathbf{x}/2, \mathbf{y}^k, \mathbf{g})$. This provides the information required to evaluate economies of scope S (as given in equation (3)).

From proposition 1, we have shown that S can be decomposed into three additive parts (see equation (4)). The parts associated with scale effects (S_R in equation (5b)) and convexity effects (S_V in equation (5c)) can be easily obtained from (6) evaluated at appropriate netput levels \mathbf{z} . The part associated with complementarity effects (S_C in (5a)) can be recovered from equation (4) by subtracting S_R and S_V from our scope measure S . This provides all the information necessary to both evaluate economies of scope S in (3) as well as its decomposition given in (4) and (5).

Finally, the analysis can be conducted with various degree of specialization. Since the complete shutdown of any operation in university production is not plausible, we will focus our attention on partial specialization scenarios where $1/2 < \beta < 1$.

VI. Results and Implications

In general, economies of scope reflect properties of the underlying technology. This means that, for a given feasible set F , any two universities using/producing similar netputs would exhibit the same economies of scope. Yet, there is much heterogeneity among universities both in terms of size and scope. In this context, a small university and a large university are located at different points of the feasible set F . Similarly, some universities are more specialized than others, which again locate them at different points of the underlying technology. Thus, given the flexible representation of the technology F allowed by estimating it non-parametrically,

economies of scope is likely to vary depending on the point of evaluation. For example, it may be that the nature of complementarities between outputs varies between small universities and large universities at different degrees of specialization. Thus, rather than compare the actual estimates, to facilitate comparisons across university types and sizes in the following scenario, we control for the degree of specialization in academic research outputs. Specifically, we first choose a point of partial integration ($\beta = 0.8$), and construct scope estimates and their decomposition for 52 universities which are on the production frontier - 36 public and 16 private – (see Appendix B, Table B-1 for a complete listing). Then, in several figures below, we explore how the level and sources of scope vary with different degrees of integration-specialization in production of patents and articles.

The economy of scope estimates associated with integrated production of academic patents and research articles are constructed using equations (3) and (6) from above. These estimates are then decomposed into their complementarity, scale, and convexity components using equations (4) and (5). Scope benefits are measured relative to faculty input: S/Fac . Given that the reference bundle \mathbf{g} represents one “unit of faculty”, this measures the proportion of faculty that can be saved by producing university outputs in a relatively integrated fashion (compared to more specialized schemes, $\beta = 1.0$). Similarly, the decompositions of scope into complementarity, scale and convexity effects are measured relative to faculty input: S_C/Fac , S_R/Fac , and S_V/Fac , respectively. From (5b’), the S_R estimate is positive, zero, or negative under IRTS, CRTS, or DRTS, respectively. We briefly discuss the patterns revealed from the 52 university estimates reported in Appendix table B-1 in order to set the context for a comparison of estimates for a representative sample of private and public universities that are depicted in Table 1 and Figures 1-4 below.

On average, scope estimates for both public and private universities are positive (0.26, 0.56), respectively. Sources of scope vary notably across university types, but several general patterns can be identified. One is that for public and private universities that are relatively small, their main source of scope comes from the scale component (See a listing of the universities in this category in Appendix B, Table B-1). This result is not surprising given that the scope simulation design divides the universities in two to compare how they would do as separate units with relatively more specialization in patents or articles. Scale is also the main source for scope estimates for the private universities, which are on average smaller than public ones. A second important pattern is that only 4 of the 52 universities show strong evidence of complementarity as a lead source of scope in this specific scenario of partial integration ($\beta = 0.8$). Finally, there is a group of eight large public universities for which the scope estimates are effectively zero, or slightly negative, because the scale component contributes negatively to their scope estimates.

This heterogeneity in the degree of scope and its sources are depicted in Table 1 and explored further in Figures 1-4 for seven public universities and six private universities. These universities were chosen on the following basis: 1/ each is on the production frontier; 2/ each is involved in significant levels of patenting; and 3/ together, they offer a cross-section representation of universities with respect to type (public vs. private), size, and extent of and sources of scope. As shown in Table 1, the proportion S/Fac varies from -0.037 (for Texas A&M University) to 1.038 (for Caltech). As in the larger sample data, the scope estimates in Table 1 reflect the finding that economies of scope are prevalent between patents and more traditional university outputs. Table 1 also illustrates the finding that the relative measures of S/Fac vary systematically across universities with these estimates being larger for smaller universities and, accordingly, for private universities.

As Table 1 illustrates, sources of scope vary substantially across universities. The relative complementarity measure S_C/Fac varies from -0.007 (for Dartmouth) to 0.264 (for MIT). The complementarity estimates are positive and small for most universities. This indicates the presence of some positive but small synergies between research publications and university patenting. For example, for the University of California-Irvine and Stanford University, the complementarity benefits amount to less than 1 percent of the faculty input. At the other extreme, the University of Wisconsin-Madison, University of Texas-Austin, and MIT exhibit relatively large complementarity benefits amounting to 11-26 percent of faculty input. This range indicates that, while complementarity benefits can be large for some universities, they are not necessarily so for all universities.

From Table 1, we can also see that some universities exhibit substantive positive economies of scope estimates but little complementarity between publication and patents (e.g., University of California-Irvine and Dartmouth). In this case, economies of scope must come from sources other than complementarity, i.e. from the scale component S_R and/or the convexity component S_V . The relative scale component S_R/Fac reported in Table 1 shows that scale effects are indeed important. In general, the larger universities exhibit a negative S_R/Fac and thus are operating in the region of decreasing returns to scale, while smaller universities are operating in the region of increasing returns to scale with positive estimates for S_R/Fac . For example, Table 1 shows that S_R/Fac varies from -0.174 (for University of Michigan) to 0.621 (for Dartmouth). This means that “being too large” (e.g., University of Michigan) can actually contribute to diseconomies of scope ($S_R < 0$). Alternatively, “being too small” (e.g., Dartmouth) contributes to economies of scope ($S_R > 0$), but perhaps in a manner that is less easily exploited (as increasing a university’s size markedly may be more difficult than adjusting the mix of outputs). In this case,

small universities (e.g., University of California-Irvine, Dartmouth) can exhibit economies of scope in the absence of complementarity because of scale effects. Additionally, some universities operating close to the region of constant returns to scale are associated with small S_R/Fac (e.g., University of California-Berkeley, University of Wisconsin-Madison). Finally, the relative convexity component S_V/Fac reported in Table 1 varies between 0.017 (for Texas A&M) and 0.546 (for Johns Hopkins). As expected, it is non-negative under a convex technology. The results indicate that the degree of convexity of the technology also varies substantially across evaluation points.

Additional estimates of relative economies of scope are presented in Figures 1 and 2. Figure 1 depicts for selected public universities how the relative scope measure S/Fac varies with the degree of specialization, β . In general, scope benefits increase with the degree of specialization, which demonstrates that the incentives of selected universities to take advantage of such benefits by combining patenting and article producing activities are most evident under scenarios associated with high degrees of specialization. However, note that this tendency also varies across universities. This increase in scope benefits is found to be modest for the University of California-Irvine, but quite large for the University of Michigan. Figure 2 depicts similar estimates for selected private universities. Figure 2 illustrates that the relative scope benefits S/Fac increase with β , strongly so for some universities (e.g., Johns Hopkins) but only mildly so for others (e.g., Dartmouth). Again, it appears that the benefits of integration across outputs depend on the degree of specialization.

Additional information on complementarity effects is presented in Figures 3 and 4 for public and private universities, respectively. These figures depict how the relative complementarity component S_C/Fac varies with the degree of specialization $\beta \in [0.5, 0.8]$. Since

$S_C = 0$ when $\beta = 0.5$, we find in general that S_C/Fac tends to increase with β . Again, this indicates that complementarity effects tend to be larger when comparing a university as an integrated firm with two more highly specialized firms. This is true for public universities as well as private universities. However, the patterns differ between public and private universities. For the former, except for the University of California-Irvine (for which changing β has little impact), S_C/Fac tends to increase significantly as β rises, reflecting the strong potential for exploiting the apparent complementarity between publications and patents by producing them in an integrated fashion. As shown in Figure 3, the rate of increase is particularly high for the University of Wisconsin-Madison for $\beta \geq 0.7$, and for the University of Texas-Austin. For these two universities, the productivity gains due to publication-patent complementarities appear to be especially large when evaluated at a level above $\beta \geq 0.7$, corresponding to a high degree of specialization. Figure 4 shows that the relative complementarity effects S_C/Fac are small for all private universities when $\beta \in [0.5, 0.65]$. Except for Johns Hopkins and MIT, they remain small for private universities (including here Cal Tech, Stanford, and Northwestern) as the degree of specialization β rises. Only above $\beta > 0.7$ for MIT does there appear to be large scope economies attributable to complementarity between articles and patents. Overall, these estimates suggest that the benefits of complementarity vary markedly across universities as well as degree of specialization, and that complementarities between articles and patents contribute significantly to scope benefits only for selected universities.

VII. Concluding Remarks

We have presented an economic analysis of scope economies at US universities, with a focus on the decomposition of economies of scope evaluated directly from the technology. We

first developed a conceptual model allowing for the investigation of economies of scope in a primal framework where the benefits of producing from an integrated firm can be measured directly from the technology of university production, using Luenberger's shortage function. This measure covers both the case of complete specialization (typically found in previous literature on economies of scope) and the case of partial specialization (suitable for investigating economies of scope in university production). Further, this approach allows for a decomposition of economies of scope into three additive parts measuring scale effects, complementarity effects and convexity effects. Relying on a non-parametric approach, we first recovered the production technology of 92 US universities using 1995-1998 data, and evaluated the associated Luenberger's shortage function. Then, measures of economies of scope and their decomposition results are obtained and analyzed.

Our analysis uncovered several important findings. First, we find that economies of scope are prevalent between patents and more traditional university outputs. Second, we documented how economies of scope measures of US universities during the 1995 to 1998 period vary with university size. We find that economies of scale (diseconomies of scale) associated with small (large) universities contribute to generating economies (diseconomies) of scope. Third, we uncovered evidence that complementarity effects are size-sensitive and vary across universities. We found large complementarity benefits between research articles and patents for a few universities, both private (MIT) and public (University of Texas-Austin, University of Wisconsin-Madison). However, such complementarity effects are found to be negligible for small universities, as well as many large universities. This suggests that synergies between articles and patents exist but are not widespread within the academic community. Fourth, our decomposition of scope effects into scale component, complementarity component and

convexity component provides useful information on the sources of scope benefits. For example, we found that scope effects tend to be important for small universities because of scale effects (and not because of complementarity effects). However, for the large public/private universities, scale effects tend to be smaller, while complementarity effects can become more important.

Our analysis suggests a need for future research to evaluate whether economies of scope may have changed over time. Also, our finding that economies of scope and complementarities can vary a lot across universities raises the question: what factors contribute to the presence of scope economies and complementarities in the research activities at U.S. universities? That undertaking appears challenging, because at the core of the university research mission is the creative process of inquiry, discovery, invention, and innovation. For example, given the complexities involved in the dynamic production of new knowledge, identifying why major complementarities in research activities arise for some universities (e.g., MIT or the University of Wisconsin-Madison) and not others may be quite difficult. Nonetheless, exploring such issues has considerable value, even if it could only identify that some complementarity and scope benefits may not be easily transferable across universities. Finally, while this paper focused on the presence and sources of scope economies at U.S. research universities, it would be useful to undertake similar analyses of other multiproduct industries (e.g., the banking industry, R&D in life sciences, the food industry, and environmental management).

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Table 1: Relative scope measure (S/Fac) and its decomposition into complementarity component (S_C/Fac), scale component (S_R/Fac), and convexity component (S_V/Fac) for selected universities (evaluated at $\beta = 0.8$).

	Expenditure on Faculty (\$1,000)	Scope (S/Fac)	Comple- mentarity (S_C/Fac)	Scale (S_R/Fac)	Convexity (S_V/Fac)
Public Universities					
UNIV MICHIGAN	138,736	0.105	0.060	-0.174	0.219
UNIV TEXAS AUSTIN	131,748	0.162	0.115	-0.007	0.054
MICHIGAN STATE UNIV	120,453	-0.051	0.059	-0.132	0.022
TEXAS A&M UNIV	109,692	-0.037	0.003	-0.057	0.017
UNIV CALIF BERKELEY	100,450	0.179	0.063	0.005	0.112
UNIV WISCONSIN MADISON	91,804	0.227	0.117	0.012	0.098
UNIV CALIF IRVINE	44,792	0.200	0.008	0.143	0.049
Private Universities					
STANFORD UNIV	80,592	0.270	0.007	0.049	0.214
MIT	80,223	0.684	0.264	0.154	0.266
NORTHWESTERN UNIV	70,883	0.109	0.040	-0.024	0.093
JOHNS HOPKINS UNIV	61,116	0.549	0.084	-0.081	0.546
CALTECH	30,181	1.038	0.011	0.594	0.433
DARTMOUTH COLL	25,823	0.647	-0.007	0.621	0.032

Figure 1: Relative Economies of Scope (S/Fac) at Selected Public Universities by Degree of Specialization (β)

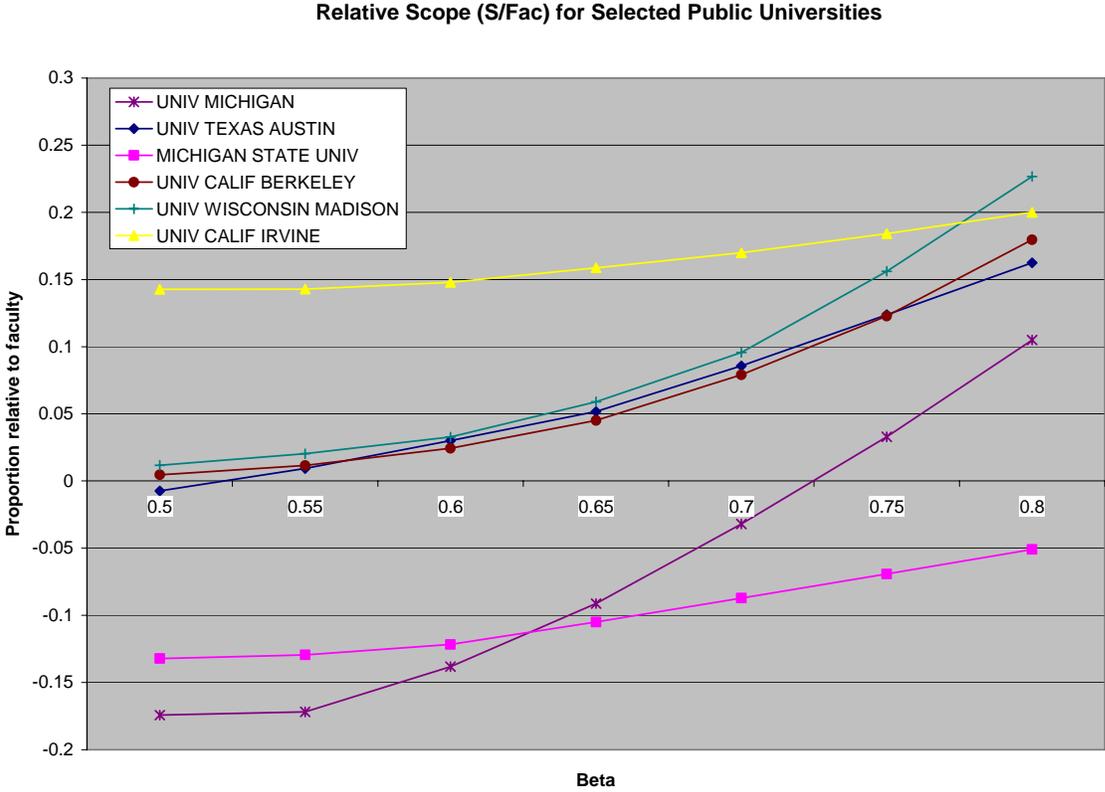


Figure 2: Relative Economies of Scope (S/Fac) at Selected Private Universities by Degree of Specialization (β)

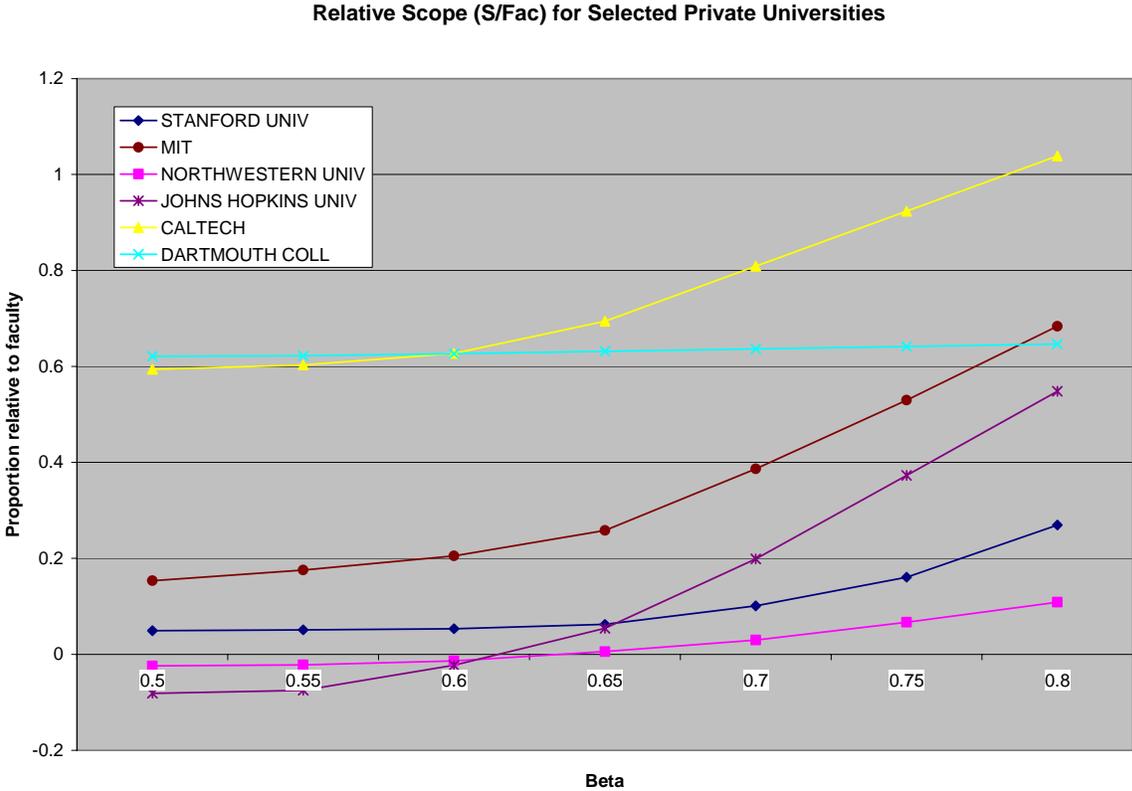


Figure 3: Relative Complementarity (S_C/Fac) at Selected Public Universities by Degree of Specialization (β)

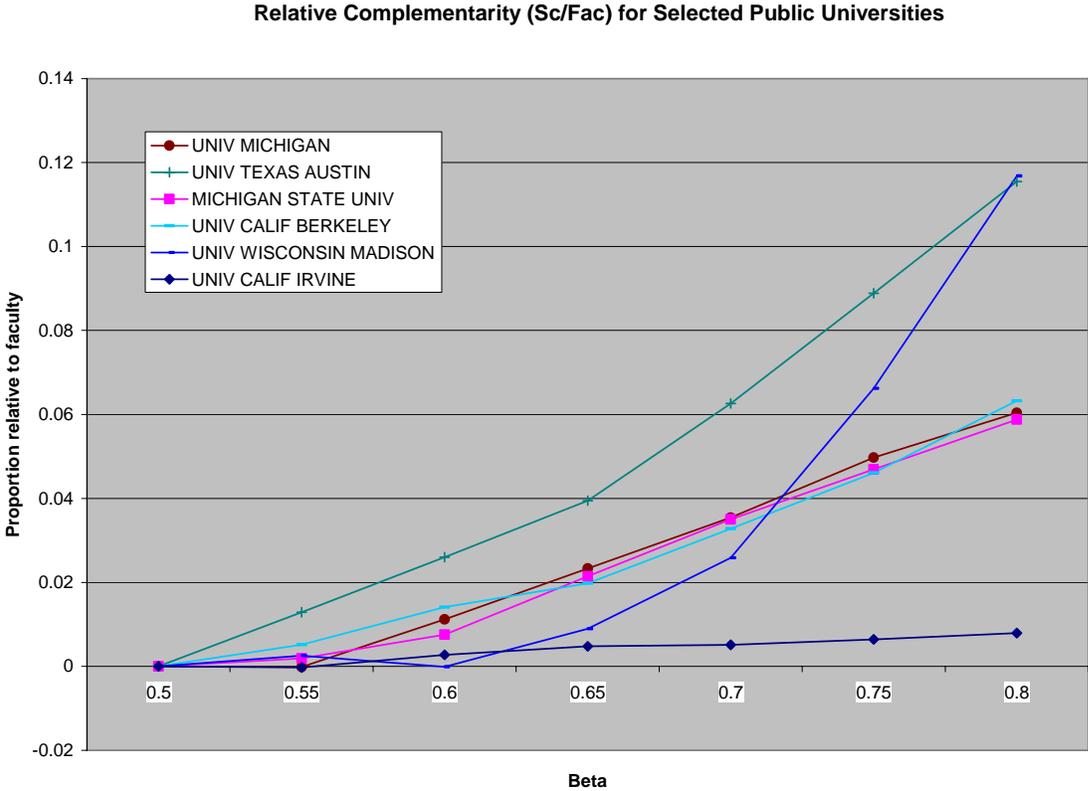
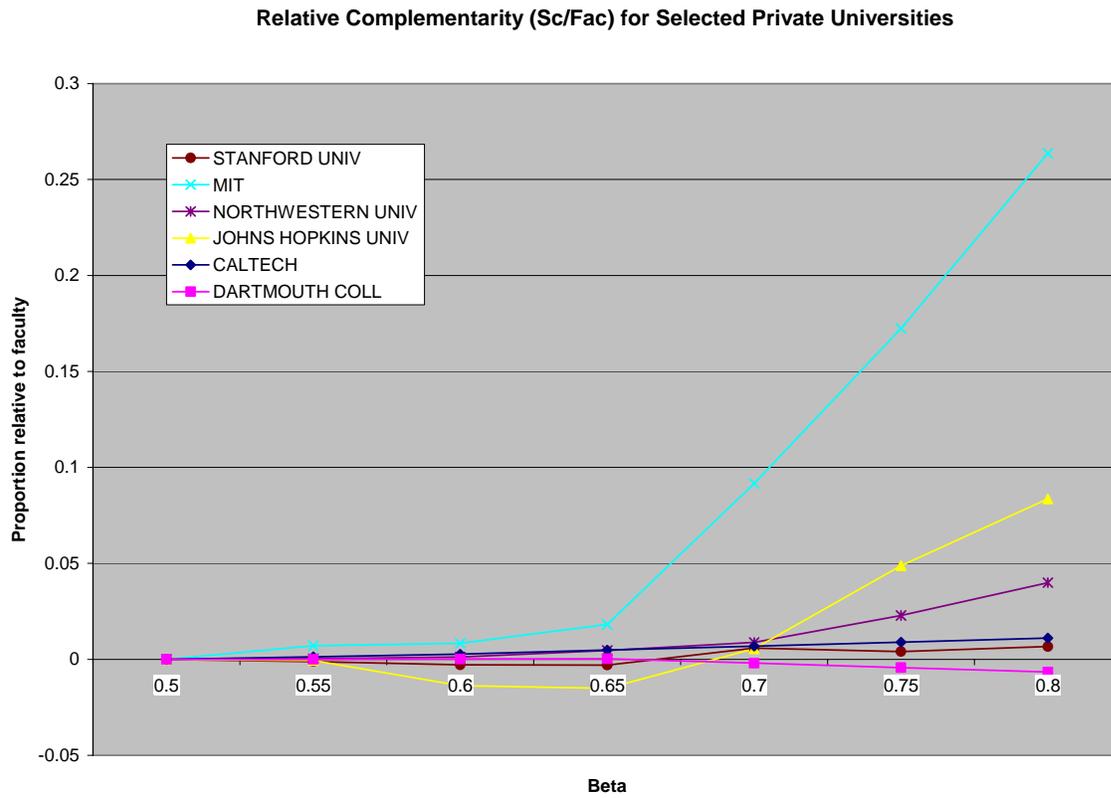


Figure 4: Relative Complementarity (S_C/Fac) for Selected Private Universities by Degree of Specialization (β)



Appendix A

Proof of Proposition 1: From equation (4), economies of scope are defined as

$$S \equiv \sigma(-\frac{1}{2} \mathbf{x}, \beta \mathbf{y}_{A1}, (1-\beta) \mathbf{y}_{A2}, \frac{1}{2} \mathbf{y}_B, \mathbf{g}) + \sigma(-\frac{1}{2} \mathbf{x}, (1-\beta) \mathbf{y}_{A1}, \beta \mathbf{y}_{A2}, \frac{1}{2} \mathbf{y}_B, \mathbf{g}) \\ - \sigma(\mathbf{x}, \mathbf{y}, \mathbf{g}) > 0.$$

When $\sigma(\mathbf{z}, \mathbf{g})$ is continuous in \mathbf{z} and differentiable almost everywhere in \mathbf{y} , this can be alternatively written as

$$S = - \int_{(1-\beta) \mathbf{y}_{A2}}^{\beta \mathbf{y}_{A2}} [\partial \sigma / \partial \gamma (-\frac{1}{2} \mathbf{x}, \beta \mathbf{y}_{A1}, \gamma, \frac{1}{2} \mathbf{y}_B, \mathbf{g}) - \partial \sigma / \partial \gamma (-\frac{1}{2} \mathbf{x}, (1-\beta) \mathbf{y}_{A1}, \gamma, \frac{1}{2} \mathbf{y}_B, \mathbf{g})] d\gamma \\ + \sigma(-\frac{1}{2} \mathbf{x}, \beta \mathbf{y}_A, \frac{1}{2} \mathbf{y}_B, \mathbf{g}) + \sigma(-\frac{1}{2} \mathbf{x}, (1-\beta) \mathbf{y}_A, \frac{1}{2} \mathbf{y}_B, \mathbf{g}) - 2 \sigma(\frac{1}{2} \mathbf{z}, \mathbf{g}), \\ + 2 \sigma(\frac{1}{2} \mathbf{z}, \mathbf{g}) - \sigma(-\mathbf{x}, \mathbf{y}, \mathbf{g}).$$

Lemma 1: $\sigma(k \mathbf{z}, \mathbf{g}) \begin{cases} < \\ = \\ > \end{cases} k \sigma(\mathbf{z}, \mathbf{g})$ under $\begin{cases} \text{DRTS} \\ \text{CRTS} \\ \text{IRTS} \end{cases}$.

Proof: By definition, the technology exhibits increasing return to scale (IRTS), constant return to scale (CRTS), or decreasing return to scale (DRTS) when, for all $\alpha > 1$, $\alpha F \subset F$, $\alpha F = F$, or $\alpha F \supset F$, respectively. Let $k \in (0, 1)$. Consider the case where there is a γ satisfying $(k \mathbf{z} - \gamma \mathbf{g}) \in F$. Then

$$\sigma(k \mathbf{z}, \mathbf{g}) = \min_{\gamma} \{ \gamma : (k \mathbf{z} - \gamma \mathbf{g}) \in F \}, \\ = k \min_{\delta} \{ \delta : (\mathbf{z} - \delta \mathbf{g}) \in (1/k) F \}, \text{ where } \delta = \gamma/k,$$

$$\begin{cases} < \\ = \\ > \end{cases} k \sigma(\mathbf{z}, \mathbf{g}) \text{ when } (1/k) F \begin{cases} \supset \\ = \\ \subset \end{cases} F, \text{ i.e., under } \begin{cases} \text{DRTS} \\ \text{CRTS} \\ \text{IRTS} \end{cases}.$$

Lemma 2: The shortage function $\sigma(\mathbf{z}, \mathbf{g})$ is convex in \mathbf{z} if F is a convex set.

Proof: Consider any two netput vectors $\mathbf{z} \in \mathbf{R}^{n+m}$ and $\mathbf{z}' \in \mathbf{R}^{n+m}$. First assume that $\sigma(\mathbf{z}, \mathbf{g})$ and

$\sigma(\mathbf{z}', \mathbf{g})$ are finite. It follows that $(\mathbf{z} - \sigma(\mathbf{z}, \mathbf{g}) \mathbf{g}) \in F$ and $(\mathbf{z}' - \sigma(\mathbf{z}', \mathbf{g}) \mathbf{g}) \in F$. Let $\mathbf{z}'' = \theta \mathbf{z}$
+ $(1-\theta) \mathbf{z}'$, for any scalar $\theta, 0 \leq \theta \leq 1$. If the set F is convex, it follows that

$$[\mathbf{z}'' - \theta \sigma(\mathbf{z}, \mathbf{g}) \mathbf{g} - (1-\theta) \sigma(\mathbf{z}', \mathbf{g}) \mathbf{g}] \in F.$$

The shortage function being defined as a minimum in (2), this yields

$$\sigma(\mathbf{z}'', \mathbf{g}) = \sigma(\theta \mathbf{z} + (1-\theta) \mathbf{z}', \mathbf{g}) \leq \theta \sigma(\mathbf{z}, \mathbf{g}) + (1-\theta) \sigma(\mathbf{z}', \mathbf{g}).$$

Second, consider the case where $\sigma(\mathbf{z}, \mathbf{g})$ and/or $\sigma(\mathbf{z}', \mathbf{g})$ are infinite. Then, the above inequality always holds. This shows that the shortage function $\sigma(\mathbf{z}, \mathbf{g})$ is convex in \mathbf{z} when F is a convex set.

Appendix B: Data

Patents

Patent data were culled from the NBER patent database, where they were identified as having a university assignee. Patents assigned to the University of California system were associated with a campus (Berkeley, Davis, Los Angeles, etc.) by the location of their authors through searches of campus directories. Relative citations for patents were generated by year and by patent class comparing each individual patent to the universe of all patents in that class (whether owned by universities or not). A university's patent count for that year is then adjusted by the ratio of number of citations received to the expected citations for that portfolio:

$$\text{Quality Adjusted Patents} = \# \text{ patents} \times \frac{\# \text{ citations received}}{E(\text{citations})}$$

where the number of expected citations, $E(\text{citations})$ is calculated as the number of citations that same portfolio of patents would receive if each patent received the average citation rate for its US patent class for that year.

Articles

Article data were culled from the ISI-Web of Science database based on universities included in their "University Science Indicators" and categories established in that same document. The Web of Science includes only the major journals in a field as identified by impact factors, such that our article measures necessarily cut out articles written for lesser journals. In addition the citation measures are only for citations in other major journals. This truncation, we believe serves our purposes of adding a subtle quality measure even to our quantity measures. Articles listed in all science disciplines were chosen.

Relative citations for articles were generated by category compared to citations of other articles assigned to the universities in the sample, rather than to all articles, and these measures were constructed annually. The same techniques of generating relative citations used for patents were used for articles.

Universities included in the sample:

Arizona State U., Boston U., Brandeis U., Brown U., Caltech, Carnegie Mellon U., Case Western Reserve U., Colorado State U., Cornell U., Dartmouth College, Emory U., Florida State

U., Georgetown U., Georgia Inst. of Technology., Harvard U., Indiana U., Iowa State U., Johns Hopkins U., Lehigh U., Loyola U., Michigan State U., MIT, N Carolina State U., New Mexico State U., Northwestern U., Ohio State U., Oregon State U., Penn State U., Princeton U., Purdue U., Rice U., Stanford U., Syracuse U., Texas A&M U., Tufts U., U. Alabama, U. Alaska, U. Arizona, U. C. Berkeley, U. C. Davis, U. C. Irvine, U. C. Los Angeles, U. C. Riverside, U. C. San Diego, U. C. Santa Barbara, U. C. Santa Cruz, U. Chicago, U. Cincinnati, U. Colorado, U. Connecticut, U. Delaware, U. Florida, U. Georgia, U. Hawaii, U. of Illinois Chicago, U. Illinois Urbana, U. Iowa, U. Kansas, U. Kentucky, U. Maryland Baltimore, U. Maryland College Park, U. Miami, U. Michigan, U. Minnesota, U. Missouri, U. N. Carolina Chapel Hill, U. Nebraska, U. New Hampshire, U. New Mexico, U. Oregon, U. Penn, U. Pittsburgh, U. Rochester, U. So Calif, U. Tennessee, U. Texas Austin, U. Texas Houston, U. Utah, U. Vermont, U. Virginia, U. Washington, U. Wisconsin Madison, Utah State U., Vanderbilt U., Virginia Polytech Inst, W. Virginia U., Wake Forest U., Washington State U., Washington U., Wayne State U., Yale U., Yeshiva U.

Table B-1: Relative scope measure (S/Fac) and its decomposition into complementarity component (S_C/Fac), scale component (S_R/Fac), and convexity component (S_V/Fac) for 52 universities (evaluated at $\beta = 0.8$).

	Scope (S/Fac)	Comple- mentarity (S_C/Fac)	Scale (S_R/Fac)	Convexity (S_V/Fac)
Private Universities				
BOSTON UNIV	0.186	0.007	0.098	0.081
BRANDEIS UNIV	1.021	0.000	1.021	0.000
CALTECH	1.038	0.011	0.594	0.433
DARTMOUTH COLL	0.647	-0.007	0.621	0.032
EMORY UNIV	0.362	-0.003	0.285	0.080
GEORGETOWN UNIV	0.245	-0.012	0.204	0.054
HARVARD UNIV	1.314	0.017	0.022	1.275
JOHNS HOPKINS UNIV	0.548	0.084	-0.081	0.546
MIT	0.684	0.264	0.154	0.266
NORTHWESTERN UNIV	0.109	0.040	-0.024	0.093
STANFORD UNIV	0.270	0.007	0.049	0.214
TUFTS UNIV	0.271	0.001	0.214	0.056
UNIV PITTSBURGH	0.076	0.023	-0.104	0.156
UNIV SO CALIF	0.147	0.001	0.127	0.019
WAKE FOREST UNIV	0.824	-0.003	0.816	0.011
YESHIVA UNIV	1.273	-0.003	1.260	0.016
AVERAGE	0.56	0.03	0.33	0.21
Public Universities				
ARIZONA STATE UNIV	-0.003	0.006	-0.011	0.002
FLORIDA STATE UNIV	0.009	0.008	-0.007	0.007
GEORGIA INST TECHNOL	0.333	-0.005	0.307	0.031
INDIANA UNIV	0.019	0.000	-0.007	0.026
MICHIGAN STATE UNIV	-0.051	0.059	-0.132	0.022
NEW MEXICO STATE UNIV	0.343	0.000	0.339	0.003
N CAROLINA STATE UNIV	0.068	0.023	0.006	0.039
OHIO STATE UNIV	-0.080	0.027	-0.162	0.054
OREGON STATE UNIV	0.477	-0.013	0.461	0.029
PENN STATE UNIV	0.094	0.051	-0.035	0.078
TEXAS A&M UNIV	-0.037	0.003	-0.057	0.017
UNIV ALABAMA	0.173	0.020	0.114	0.039
UNIV COLORADO	0.134	-0.001	0.103	0.032
UNIV FLORIDA	0.011	0.041	-0.095	0.065
UNIV ILLINOIS URBANA	-0.040	0.032	-0.177	0.104
UNIV MICHIGAN	0.105	0.060	-0.174	0.219
UNIV MINNESOTA	0.112	0.026	-0.104	0.191
UNIV NEW HAMPSHIRE	0.459	0.000	0.459	0.000
UNIV OREGON	0.381	0.000	0.380	0.001
UNIV TENNESSEE	0.092	0.011	0.014	0.067

UNIV TEXAS HOUSTON	1.242	0.000	1.203	0.039
UNIV TEXAS AUSTIN	0.162	0.115	-0.007	0.054
UNIV UTAH	0.210	0.026	0.139	0.045
UNIV WASHINGTON	0.318	0.102	-0.024	0.240
UNIV WISCONSIN MADISON	0.227	0.117	0.012	0.098
UNIV ALASKA	1.458	0.000	1.458	0.000
UTAH STATE UNIV	0.366	0.000	0.364	0.002
UNIV VERMONT	0.566	0.000	0.561	0.005
W VIRGINIA UNIV	0.219	0.000	0.219	0.000
UNIV CALIF BERKELEY	0.180	0.063	0.005	0.112
UNIV CALIF DAVIS	0.168	0.015	0.081	0.072
UNIV CALIF IRVINE	0.200	0.008	0.143	0.049
UNIV CALIF LOS ANGELES	0.245	0.047	0.010	0.188
UNIV CALIF SAN DIEGO	0.386	0.036	0.149	0.201
UNIV CALIF SANTA BARBARA	0.222	-0.012	0.175	0.059
UNIV CALIF SANTA CRUZ	0.604	0.000	0.601	0.003
AVERAGE	0.26	0.02	0.18	0.06

* 9 private universities with dominant scale components: BRANDEIS UNIV, CALTECH, DARTMOUTH COLL, EMORY UNIV, GEORGETOWN UNIV, TUFTS UNIV, UNIV SO CALIF, WAKE FOREST UNIV, YESHIVA UNIV

* 15 public universities with dominant scale components: GEORGIA INST TECHNOL, NEW MEXICO STATE UNIV, OREGON STATE UNIV, UNIV ALABAMA, UNIV COLORADO, UNIV NEW HAMPSHIRE, UNIV OREGON, UNIV TEXAS HOUSTON, UNIV UTAH, UNIV ALASKA, UTAH STATE UNIV, UNIV VERMONT, W VIRGINIA UNIV, UNIV CALIF IRVINE, UNIV CALIF SANTA BARBARA, UNIV CALIF SANTA CRUZ

* 4 universities with strong evidence of complementarity as a source of scope: MIT, UNIV TEXAS AUSTIN, UNIV WASHINGTON, UNIV WISCONSIN MADISON

* 8 larger public universities with zero or negative scope estimates: ARIZONA STATE UNIV, FLORIDA STATE UNIV, INDIANA UNIV, MICHIGAN STATE UNIV, N CAROLINA STATE UNIV, TEXAS A&M UNIV, UNIV FLORIDA, UNIV ILLINOIS URBANA

Footnotes

¹ Note that other measures have been developed in the literature. They include the directional distance function $D_g(\mathbf{z}, \mathbf{g})$ discussed by Robert G. Chambers, Yangho Chung, and Rolf Färe and Rolf Färe, and Shawna Grosskopf: $D_g(\mathbf{z}, \mathbf{g}) \equiv \max_{\beta} \{\beta: (\mathbf{z} + \beta \mathbf{g}) \in F\}$. Since it satisfies $\sigma(\mathbf{z}, \mathbf{g}) = -D_g(\mathbf{z}, \mathbf{g})$, it should be clear that the analysis presented below could be presented equivalently using the directional distance function. Other measures include Shephard's output distance function $D_O(\mathbf{z}) \equiv \min_{\theta} \{\theta: (-\mathbf{x}, \mathbf{y}/\theta) \in F\}$, and Shephard's input distance function $D_I(\mathbf{z}) \equiv \max_{\theta} \{\theta: (-\mathbf{x}/\theta, \mathbf{y}) \in F\}$. The relationships between these functions and the shortage function have been analyzed in the literature (Rolf Färe, and Shawna Grosskopf; Robert G. Chambers, Yangho Chung, and Rolf Färe). However, by measuring input or output proportions, the Shephard's functions are not additive across firms. As such they do not provide attractive measurements for analyzing economies of scope.

² Note that, in the case where $I_B = \emptyset$ and $\beta = 1$, this involves no loss of generality since any partition of I_A can always be decomposed into a series of binary partitions.

³ Below, the linear programming problem (6) is solved using GAMS software.